Cardiovascular safety of exenatide BID: an integrated analysis from controlled clinical trials in participants with type 2 diabetes

Robert Ratner1, Jenny Han2, Dawn Nicewarner2, Irina Yushmanova2, Byron J Hoogwerf3, Larry Shen2*

Abstract
It is important for patients that treatments for diabetes not increase cardiovascular (CV) risk. The objective of this analysis was to examine retrospectively the CV safety of exenatide BID, a GLP-1 receptor agonist approved for treating hyperglycemia in patients with type 2 diabetes not adequately controlled with diet and exercise. Individual participant data was pooled to assess the relative risk (RR) of CV events with exenatide BID versus a pooled comparator (PC) group treated with either placebo or insulin from 12 controlled, randomized, clinical trials ranging from 12-52 weeks. Mean baseline values for HbA1c (8.33-8.38%), BMI (31.3-31.5 kg/m²), and duration of diabetes (8 y) were similar between groups. Trials included patients with histories of microvascular and/or macrovascular disease. Customized primary major adverse CV events (MACE) included stroke, myocardial infarction, cardiac mortality, acute coronary syndrome, and revascularization procedures. The Primary MACE RR (0.7; 95% CI 0.38, 1.31), calculated by the Mantel-Haenszel method (stratified by study), suggested that exenatide use (vs. PC) did not increase CV risk; this result was consistent across multiple analytic methods. Because the trials were not designed to assess CV outcomes, events were identified retrospectively from a list of preferred terms by physicians blinded to treatment. Other limitations included the low number of CV events, the short duration of trials (<1 y), and a single active comparator (insulin). The results of these analyses are consistent with those of a recent retrospective analysis of a large insurance database that found that patients treated with exenatide twice daily were less likely to have a CV event than were patients treated with other glucose-lowering therapies.

Keywords: GLP-1 receptor agonist, diabetes, cardiovascular safety

Background
Despite the numerous advancements in glucose-lowering medications in recent years, the incidence of cardiovascular (CV) morbidity and mortality in patients with type 2 diabetes has not consistently decreased [1]. In 2007, a controversial, widely publicized meta-analysis of 42 trials suggested that rosiglitazone, a thiazolidinedione (TZD), was associated with increased risk of myocardial infarction (MI) and death due to CV events [2-6]. At that time, a joint advisory committee meeting concluded that rosiglitazone increased myocardial ischemia and the FDA added a black box warning to the rosiglitazone label [2,7]. Although a non-inferiority CV outcomes trial, RECORD concluded that rosiglitazone did not increase CV morbidity or mortality compared with other glucose-lowering medications [8]; the open-label and unblinded design of RECORD prompted some to question the quality of the data [7]. Following updated meta-analyses and recent congressional inquiry, concerns arose again about the CV safety of rosiglitazone [9-12]. Recently, an advisory committee meeting (July 2010) concluded that rosiglitazone significantly increased CV risk [7].

The risk-to-benefit profile of medications may begin to take on greater importance for future approvals of new drugs [13,14]. Exenatide twice daily (exenatide BID) was the first GLP-1 receptor agonist approved for treatment of hyperglycemia in patients with type 2 diabetes not adequately controlled with diet and exercise, sulfonylurea (SU), TZD, or metformin (MET; alone, or with an SU, or TZD). In addition to glycemic control and a low risk of hypoglycemia, exenatide may favorably affect

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several CV risk factors, such as blood pressure, lipid profiles, and body weight [15-17]. Because the incidence of CV events was low in each of the clinical trials, we undertook this pooled analysis to provide an integrated assessment of the CV safety data from 12 studies with exenatide BID.

Methods

Study Selection

For each clinical trial, a protocol was approved by an institutional review board, in accordance with the principles described in the Declaration of Helsinki (World Medical Association 1997). Each of the trials included in this analysis was conducted by Amylin Pharmaceuticals, Inc. and Eli Lilly and Company. Complete efficacy and safety data for these trials have been reported previously [18-29]. Randomized, controlled trials (8 blinded, 4 open-label) that were completed by September 30, 2008 and of at least 12 weeks in duration, that were included in this analysis were designed to compare the efficacy and tolerability of exenatide BID to placebo or insulin (Table 1). They were either placebo-controlled or active-controlled studies where insulin served as the comparator. The studies included in this analysis were not designed to assess CV events and events were not prospectively adjudicated.

Analysis population

The current analyses included patients from the intent-to-treat populations (i.e., patients who received at least 1 dose of randomized study medication) of each study (Table 1). Patients in all studies had type 2 diabetes and were treated continuously with exenatide and MET, SFU, or TZD alone or in combination. All patients were 18 to 75 years-of-age, had a HbA1c ≤11.0%, a body mass index (BMI) of 25 to 45 kg/m², and a history of stable body weight (≤10% change) for at least 3 months. Patients were excluded if they had used weight loss drugs or had evidence of a significant medical condition. Investigators were asked to exclude patients with evidence of active cardiac disease within 1 year prior to the study, i.e., MI, clinically significant arrhythmia, unstable angina, moderate to severe congestive heart failure, coronary artery bypass surgery, or angioplasty. Patients with a >1 year history of MI, transient ischemic attack or large vessel disease or with a history of microvascular disease were eligible for enrollment. All patients provided written informed consent before participation.

Outcomes

Events were identified by preferred terms according to the Medical Dictionary for Regulatory Activities (MedDRA 11.0). A team of physicians, blinded to treatment, independently reviewed the list of preferred terms prior to the analyses to focus on the terms most likely to represent true events of interest, regardless of occurrence. Blinded adjudication of the CV events was not pre-specified in the study protocols; therefore, events were identified retrospectively using the pre-specified list of preferred terms independent of whether or not the events had occurred. At the time of the occurrence during the original trial, the events

Table 1 Summary of controlled studies included in CV analysis

<table>
<thead>
<tr>
<th>Study/Registry Number</th>
<th>Diabetes Management</th>
<th>Duration* (Weeks)</th>
<th>ITT (N)</th>
<th>Exposure (SY)</th>
<th>Placebo or Comparator</th>
</tr>
</thead>
<tbody>
<tr>
<td>DeFronzo et al, 200518</td>
<td>Met</td>
<td>30</td>
<td>223</td>
<td>113.8</td>
<td>113 57.8</td>
</tr>
<tr>
<td>Buse et al, 200420</td>
<td>SFU</td>
<td>30</td>
<td>254</td>
<td>123.2</td>
<td>123 55.1</td>
</tr>
<tr>
<td>Kendall et al, 200519</td>
<td>Met + SFU</td>
<td>30</td>
<td>486</td>
<td>254.9</td>
<td>247 122.2</td>
</tr>
<tr>
<td>Zinman et al, 200724</td>
<td>TZD ± Met</td>
<td>16</td>
<td>121</td>
<td>31.7</td>
<td>112 32.3</td>
</tr>
<tr>
<td>Kadowaki et al, 200827</td>
<td>SFU ± Met</td>
<td>12</td>
<td>111</td>
<td>23.9</td>
<td>40 9.2</td>
</tr>
<tr>
<td>Gao, et al, 200926</td>
<td>Met ± SFU</td>
<td>16</td>
<td>234</td>
<td>65.5</td>
<td>233 67.3</td>
</tr>
<tr>
<td>Moretto et al, 200828</td>
<td>D + E</td>
<td>24</td>
<td>155</td>
<td>65.2</td>
<td>77 33.1</td>
</tr>
<tr>
<td>Gill et al, 201029</td>
<td>Met and/or TZD</td>
<td>12</td>
<td>28</td>
<td>5.8</td>
<td>26 5.7</td>
</tr>
<tr>
<td>Heine et al, 200522</td>
<td>Met + SFU</td>
<td>26</td>
<td>282</td>
<td>122.5</td>
<td>267 124.6</td>
</tr>
<tr>
<td>Nauck et al, 200723</td>
<td>Met + SFU</td>
<td>52</td>
<td>52</td>
<td>220.1</td>
<td>248 228.6</td>
</tr>
<tr>
<td>Davis et al, 200725</td>
<td>SFU or Meg and/or Met</td>
<td>16</td>
<td>33</td>
<td>7.7</td>
<td>16 5.2</td>
</tr>
<tr>
<td>Barnett et al, 200721</td>
<td>Met or SFU</td>
<td>16‡</td>
<td>136</td>
<td>37.3</td>
<td>127 38.9</td>
</tr>
<tr>
<td>Totals</td>
<td>–</td>
<td>–</td>
<td>2,316</td>
<td>1,071.6</td>
<td>1,629 779.9</td>
</tr>
</tbody>
</table>

D + E = diet and exercise therapy; ITT = Intent-to-Treat Population; Meg = meglitinide; Met = metformin; OAD = oral antidiabetic medications; SFU = sulfonylurea; TZD = thiazolidinedione.

Duration of treatment with randomized study medication.

Includes treatment with exenatide 2.5 mcg or 5 mcg BID for duration of study, or 4 weeks of exenatide 5 mcg BID followed by exenatide 10 mcg BID for remainder of study.

NCT00099619 had a crossover design, with 16 weeks per period (exenatide or insulin glargine).
were closely reviewed in accordance with normal clinical trial monitoring and follow-up of adverse events, laboratory evaluations, physical examinations, vital signs measurements, with particular attention to serious adverse events.

All data for patients who died were examined to ascertain if the underlying cause was CV in nature based on the preferred term provided and the cases were reviewed in detail. “Sudden deaths” were adjudicated as CV events, in conformance with most large CV outcomes trials [30-35]. The final list of terms was concordant with the FDA list presented at the advisory committee meeting in April 2009 for other antidiabetic agents [36].

Primary Outcome
The primary outcome was Primary Major Adverse CV Events (MACE); per FDA guidance, it included terms reflective of CV mortality, stroke, myocardial infarction, acute coronary syndrome, and revascularization procedures.

Secondary Outcome
The secondary CV endpoint included all relevant CV adverse events. This expanded endpoint comprised all terms of the Primary MACE endpoint plus terms for arrhythmia, heart failure (with or without hospitalization), and mechanical-related events. Mechanical-related events were aortic valve disease, aortic valve stenosis, cardiac failure congestive, cardiomegaly, CV disorder, heart valve incompetence, left atrial dilatation, mitral valve incompetence, and tricuspid valve incompetence. Heart failure and mechanical-related adverse events were included in an effort to encompass as many potentially important CV-related events as possible.

Analysis
A meta-analysis was performed on 12 completed longer-term (3- to 12-month), randomized, placebo- or insulin comparator-controlled trials of exenatide BID, in accordance with the FDA guidelines [37]. Data from approximately 4,000 patients with type 2 diabetes and an average exposure of 24 weeks were included [18-29]. Pooled data from placebo- and insulin-treated patients were compared with pooled data from exenatide-treated patients. The exenatide cohort included participants randomized to receive exenatide BID 2.5 mcg (n = 37), 5 mcg (n = 594), or 10 mcg (n = 1,685). Because of the low number of CV events, the 3 exenatide dose groups were combined.

For the primary analysis, the Relative Risk (RR) of an incident CV event and the corresponding 95% confidence interval (CI) were calculated using Mantel-Haenszel method stratified by study. In order to demonstrate robustness, the RR and its 95% CI were calculated using the following additional methods: pooled RR (without stratification by study and with common continuity correction [i.e., adding 0.5 to all cells if one of the treatment groups had no events]) and Shuster’s RR [38] (weights all trials equally and included studies with no events). The Hazard Ratio (HR) was calculated using 2 methods: 1) the Cox proportional hazard model (time to first event) with adjustment for study, and 2) the Andersen-Gill model (recurrent events) with adjustment for study. Ninety-five percent CIs for the RR and HR were provided at a 2-sided significance level of 0.05. Weighted Kaplan-Meier survival curves were generated to show the time to first event and the proportion of patients who were risk-free over time [39]. The time to event was calculated from the first randomized dose to the time of the first cardiac event. Exposure Adjusted Incidence Rate (EAIR) and its 95% CI were calculated using the Exact method [40]. The RR and its 95% CI based on EAIR and event rate were provided [41].

Additional subgroup analyses by age (<65 vs. ≥65), BMI (<30 vs. ≥30), and renal function (normal, mild impairment, and moderate impairment) were provided to assess the effects of these baseline characteristics. SAS 9.2® (Statistical Analysis Software, Cary, NC, USA) was used for all analyses.

Results
The data included in these analyses represented 1,072 patient-years (PY) exposure with exenatide BID (N = 2,316) and 780 PY exposure with comparators (placebo, n = 971; insulin, n = 658). Demographic and baseline characteristics were similar between treatment groups (Table 2). Participants had similar mean baseline values for HbA1c (8.33%-8.38%), BMI (31.3-31.5 kg/m²), and duration of diabetes (8 years). Of note, the mean systolic blood pressures at baseline for both cohorts were 131 ± 4 mm Hg and 132 ± 5 mm Hg for the exenatide and the pooled comparator cohorts, respectively. Some patients had a past history of CV disease (including multiple events) and/or microvascular disease (including multiple conditions). From baseline to endpoint, heart rate changes were +0.5 ± 9.8 beats per minute (bpm) for exenatide and +0.1 ± 9.2 bpm for the pooled comparators (mean ± SD).

Incidence rates of Primary MACE were similar for 5 mcg (0.9%) and 10 mcg (1.1%) exenatide dose groups, and no events were observed in the 2.5 mcg group. Therefore, data from all exenatide doses (2.5 mcg, 5 mcg, 10 mcg) were pooled in the analyses. The EAIR for exenatide treatment group was 18.73 per 1,000 patient-years versus 23.17 per 1,000 patient-years for the pooled comparator group (Table 3). Overall, 26 participants experienced serious adverse CV events: 15 (0.6%)
in the exenatide group and 11 (0.7%) in the pooled comparator group (data not shown). Four fatal CV events occurred: 2 in the exenatide group (MI, atrial fibrillation) and 2 in the comparator group (MI, cerebrovascular accident).

The point estimates for Primary MACE and Secondary CV endpoints between exenatide BID and the pooled comparator were both 0.7 (the upper limits of the 95% CIs were 1.3 and 1.0, respectively), in favor of exenatide (Figure 1). These results suggest that exenatide did not increase the CV risk. For the individual studies, the point estimates for RR were <1.0 (favoring exenatide) for 10 of 12 long-term controlled studies (Figure 2). In addition, the RRs and 95% CIs for the Primary MACE were consistent across multiple methods of analysis (including sensitivity analyses of RR based on incidence and RR based on EAIR event rate), with point estimates ranging from 0.5 to 0.8, and the upper limit of the 95% CIs ranging from 1.3 to 1.5 (Figure 3). Similar point estimates were observed for the broader secondary CV endpoint; however, the upper limit of the 95% CI was <1.3 for all methods (Figure 3). A weighted Kaplan-Meier plot shows that a significantly higher percentage of exenatide-treated patients than pooled comparator-treated patients were free of a primary MACE event over 1 year (*P* < 0.0001; Figure 4).

### Discussion

**Assessing the CV Risk of Glucose-Lowering Therapies**

One of the strengths of this meta-analysis is the use of individual participant data from each of the trials, in contrast with a typical meta-analysis in which summary statistics for individual studies are used. Another strength was the use of multiple methods of analysis, some of which allowed for the inclusion of studies with zero events in the exenatide group and 11 (0.7%) in the pooled comparator group (data not shown). Four fatal CV events occurred: 2 in the exenatide group (MI, atrial fibrillation) and 2 in the comparator group (MI, cerebrovascular accident).

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### Table 2 Baseline characteristics and demographics

<table>
<thead>
<tr>
<th>Baseline Characteristics</th>
<th>Exenatide (N = 2,316)</th>
<th>Pooled Comparator (N = 1,629)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Gender, M/F (%)</td>
<td>56/44</td>
<td>53/47</td>
</tr>
<tr>
<td>Race (%)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Caucasian</td>
<td>64</td>
<td>65</td>
</tr>
<tr>
<td>Black</td>
<td>6</td>
<td>4</td>
</tr>
<tr>
<td>Hispanic</td>
<td>11</td>
<td>10</td>
</tr>
<tr>
<td>Asian</td>
<td>19</td>
<td>20</td>
</tr>
<tr>
<td>Other</td>
<td>1</td>
<td>1</td>
</tr>
<tr>
<td>Age (years)</td>
<td>56 ± 10</td>
<td>56 ± 10</td>
</tr>
<tr>
<td>≥65 years (%)</td>
<td>21</td>
<td>20</td>
</tr>
<tr>
<td>Duration of Diabetes (years)</td>
<td>8 ± 6</td>
<td>8 ± 6</td>
</tr>
<tr>
<td>≥10 years (%)</td>
<td>32</td>
<td>35</td>
</tr>
<tr>
<td>Weight (kg)</td>
<td>89.1 ± 20.3</td>
<td>87.8 ± 19.8</td>
</tr>
<tr>
<td>Body Mass Index (BMI [kg/m²])</td>
<td>31.5 ± 5.6</td>
<td>31.3 ± 5.4</td>
</tr>
<tr>
<td>BMI &lt;30 kg/m² (%)</td>
<td>44</td>
<td>44</td>
</tr>
<tr>
<td>BMI ≥30 kg/m² (%)</td>
<td>56</td>
<td>55</td>
</tr>
<tr>
<td>HbA1c (%)</td>
<td>8.33 ± 1.06</td>
<td>8.38 ± 1.07</td>
</tr>
<tr>
<td>Systolic Blood Pressure (SBP [mm Hg])</td>
<td>131 ± 4</td>
<td>132 ± 5</td>
</tr>
<tr>
<td>Diastolic Blood Pressure (DBP [mm Hg])</td>
<td>79 ± 2</td>
<td>79 ± 1</td>
</tr>
<tr>
<td>Heart Rate (beats per minute)</td>
<td>75 ± 9</td>
<td>75 ± 9</td>
</tr>
<tr>
<td>Renal Function Impairment* (%)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>None</td>
<td>86</td>
<td>86</td>
</tr>
<tr>
<td>Mild</td>
<td>13</td>
<td>14</td>
</tr>
<tr>
<td>Moderate</td>
<td>1</td>
<td>1</td>
</tr>
</tbody>
</table>

Data are mean ± SD unless otherwise noted. Percentages may not add to 100 due to rounding.

*Renal function is defined based on creatinine clearance (CrCl) as calculated by the Cockcroft-Gault equation: normal, CrCl >80 mL/min; mild impairment, CrCl >50-80 mL/min; moderate impairment, CrCl >30-50 mL/min.

### Table 3 Incidence rates and exposure-adjusted incidence rates for primary MACE and secondary CV endpoints

<table>
<thead>
<tr>
<th></th>
<th>Primary MACE Endpoint</th>
<th>Secondary MACE Endpoint</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Exenatide (N = 2,316)</td>
<td>Pooled Comparator (N = 1,629)</td>
</tr>
<tr>
<td><strong>Primary Analyses</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incidence (n)</td>
<td>20</td>
<td>18</td>
</tr>
<tr>
<td>Incidence (n/N)</td>
<td>0.009</td>
<td>0.011</td>
</tr>
<tr>
<td>RR (95% CI)</td>
<td>0.70 (0.38, 1.31)</td>
<td></td>
</tr>
<tr>
<td><strong>Secondary Analysis</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>EAIR (per 1,000 years)</td>
<td>18.73</td>
<td>23.17</td>
</tr>
<tr>
<td>RR (95% CI)</td>
<td>0.81 (0.43, 1.53)</td>
<td></td>
</tr>
<tr>
<td>Event Rate (per 1,000 years)</td>
<td>22.40</td>
<td>28.21</td>
</tr>
<tr>
<td>RR (95% CI)</td>
<td>0.79 (0.45, 1.42)</td>
<td></td>
</tr>
</tbody>
</table>

CI = confidence interval; CV = cardiovascular; EAIR = exposure-adjusted incidence rate; MACE = major adverse cardiovascular events; RR = risk ratio.

RR for the incidence was calculated using the Mantel-Haenszel estimate with study as a stratification factor.

EAIR and event rate were calculated based on the Exact method, with corresponding RR calculated using the log-normal approximation.
CV event rates. The results of this meta-analysis provide a preliminary assessment of the CV risk associated with exenatide treatment across trials used for registration. Because these findings are not conclusive, a prospective, adequately-powered, adjudicated, CV outcome study (NCT01144338) of the investigational once-weekly formulation of exenatide was initiated.

Applying the FDA guidance on how to assess CV safety in the development of treatments for type 2 diabetes to this analysis, the HR point estimates were both <1 for the Primary MACE and Secondary CV endpoints resulting from the lower observed incidence of CV events with exenatide BID compared to the control group [37]. These results statistically excluded a 1.8-fold increase of the CV risk by exenatide and were consistent with the FDA guideline that state that the upper limit of a 2-sided 95% CI for new diabetes therapies should be <1.8. In fact, the results suggest that exenatide treatment may improve CV outcomes, although this effect failed to reach statistical significance. Our results are consistent with those of a recent retrospective analysis of a large insurance database in which the 39,275 patients treated with exenatide twice daily were found to be significantly less likely to have a CV event than were the 381,218 patients treated with other glucose-lowering therapies (HR 0.81, 95% CI 0.68-0.91, P = 0.01) [42]. Although the results of the current meta-analysis suggest that there may be an association between exenatide and improved CV outcomes, data from multi-year, randomized, controlled, adequately-powered clinical trials with prospective blinded adjudication of CV events are

<table>
<thead>
<tr>
<th>Study</th>
<th>Exenatide n/N (%)</th>
<th>Placebo or Comparator n/N (%)</th>
<th>RR</th>
</tr>
</thead>
<tbody>
<tr>
<td>Defronzo et al, 2005</td>
<td>1/223 (0.45)</td>
<td>5/113 (4.42)</td>
<td>0.10</td>
</tr>
<tr>
<td>Buse et al, 2004</td>
<td>4/254 (1.57)</td>
<td>1/123 (0.81)</td>
<td>1.94</td>
</tr>
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<td>Kendall et al, 2005</td>
<td>8/486 (1.65)</td>
<td>5/247 (2.02)</td>
<td>0.81</td>
</tr>
<tr>
<td>Kadowaki et al, 2009</td>
<td>0/111 (0.00)</td>
<td>1/40 (2.50)</td>
<td>0.12</td>
</tr>
<tr>
<td>Zinman et al, 2007</td>
<td>0/121 (0.00)</td>
<td>0/112 (0.00)</td>
<td>0.93</td>
</tr>
<tr>
<td>Gao, et al, 2009</td>
<td>1/234 (0.43)</td>
<td>2/233 (0.86)</td>
<td>0.50</td>
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<tr>
<td>Moretto et al, 2008</td>
<td>0/155 (0.00)</td>
<td>0/77 (0.00)</td>
<td>0.50</td>
</tr>
<tr>
<td>Gill et al, 2010</td>
<td>0/28 (0.00)</td>
<td>0/26 (0.00)</td>
<td>0.93</td>
</tr>
<tr>
<td>All Placebo controlled</td>
<td>14/1612 (0.01)</td>
<td>14/971 (0.01)</td>
<td>0.53</td>
</tr>
<tr>
<td>Heine et al, 2005</td>
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<td>1/267 (0.37)</td>
<td>0.95</td>
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<td>Nauck et al, 2007</td>
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<td>3/248 (1.21)</td>
<td>1.63</td>
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<td>Davis et al, 2007</td>
<td>0/33 (0.00)</td>
<td>0/16 (0.00)</td>
<td>0.50</td>
</tr>
<tr>
<td>Barnett et al, 2007</td>
<td>0/136 (0.00)</td>
<td>0/127 (0.00)</td>
<td>0.93</td>
</tr>
<tr>
<td>All Comparator controlled</td>
<td>8/704 (0.01)</td>
<td>4/658 (0.01)</td>
<td>1.46</td>
</tr>
<tr>
<td>Totals</td>
<td>20/2316 (0.86)</td>
<td>18/1629 (1.10)</td>
<td>0.70</td>
</tr>
</tbody>
</table>

* RR exenatide BID vs placebo or comparator by individual study; RR exenatide BID vs All placebo studies; RR exenatide vs All comparator studies; RR exenatide BID vs All placebo or comparator studies.

Figure 1 Risk of primary MACE and secondary CV endpoints with exenatide BID relative to pooled comparators.

Figure 2 Incidence and RR for primary MACE with exenatide use by individual study.
needed to evaluate whether exenatide has cardio-protective effects.

Despite the increased relative risk of CV disease in diabetes, the absolute risk may only be reduced when multifactorial treatment strategies are used [43-45]; therefore, the CV risk associated with glucose-lowering agents is not easily ascertained. Even in randomized, controlled trials that included CV outcomes as part of the endpoint, the outcomes may be difficult to interpret.

The University Group Diabetes Program (UGDP) suggested treatment with tolbutamide, a first generation SUF, was associated with increased risk of negative CV outcomes [46]. Despite the limitations of UGDP, the implication that SFUs were associated with some risk remained until the results of the landmark UKPDS were available [47]. From UKPDS, it was evident that SFU or MET monotherapy was associated with reduced CV risk; however, treatment with the combination of SFU and MET was associated with increased CV risk [48]. Demonstration of the impact of glycemic control on the risk of acute MI was limited by the relatively low absolute MI event rate, but long-term follow-up studies of UKPDS and DCCT found that intense glycemic control reduced major CV events [31,49,50].

Four large trials have analyzed the effects of glucose lowering on CV risk [30,31,35,47,49,51]. The UKPDS studied patients early in the course of disease and found that intensive therapy was associated with a reduction in CV events, especially in the small cohort of obese patients who were treated with metformin [47,49]. The 3 large trials (ADVANCE, ACCORD, and VADT) in patients with longer duration of disease generally had HRs less than 1.0 with intensive therapy, but none achieved statistical significance [30,35,51]. The systematic review of these 4 trials in a pooled analysis did show a reduction in CV disease (RR 0.90, 95% CI

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### Endpoint/ Method

<table>
<thead>
<tr>
<th>Endpoint/ Method</th>
<th>RR</th>
<th>(95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Primary MACE</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>RR (Mantel-Haenszel)</td>
<td>0.70</td>
<td>(0.38, 1.31)</td>
</tr>
<tr>
<td>HR (Cox)</td>
<td>0.71</td>
<td>(0.36, 1.37)</td>
</tr>
<tr>
<td>HR (Andersen-Gill)</td>
<td>0.69</td>
<td>(0.39, 1.25)</td>
</tr>
<tr>
<td>RR (Pooled)</td>
<td>0.78</td>
<td>(0.42, 1.47)</td>
</tr>
<tr>
<td>RR (Shuster)</td>
<td>0.53</td>
<td>(0.21, 1.35)</td>
</tr>
<tr>
<td><strong>Secondary CV endpoint</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>RR (Mantel-Haenszel)</td>
<td>0.69</td>
<td>(0.46, 1.03)</td>
</tr>
<tr>
<td>HR (Cox)</td>
<td>0.68</td>
<td>(0.44, 1.04)</td>
</tr>
<tr>
<td>HR (Andersen-Gill)</td>
<td>0.69</td>
<td>(0.47, 1.01)</td>
</tr>
<tr>
<td>RR (Pooled)</td>
<td>0.77</td>
<td>(0.51, 1.17)</td>
</tr>
<tr>
<td>RR (Shuster)</td>
<td>0.44</td>
<td>(0.22, 0.86)</td>
</tr>
</tbody>
</table>

* RR exenatide BID vs. all placebo or comparator studies

Sensitivity analyses included the pooled RR without stratification by study, and Shuster’s method, which weights all trials equally. Hazard ratios (HRs) and 95% CIs were obtained from the Cox proportional hazard regression model and the Andersen-Gill model, with adjustment for study.

Figure 3 Forest plot of CV end points by statistical method.

Figure 4 Weighted Kaplan Meier plot for subjects without a primary MACE event by treatment in controlled studies of exenatide BID.
0.83-0.98) and coronary heart disease (RR 0.89, 95% CI 0.81-0.96), but not in stroke, coronary heart failure, CV disease mortality, or all-cause mortality [52]. ACCORD and VADT reported that mortality was associated with hypoglycemia [35,51,53]. In fact, ACCORD was discontinued because of this increased mortality. Hypoglycemia in each of the arms of the ACCORD trial was associated with increased mortality, though hypoglycemia did not explain the increased mortality risk in the intensive arm [53]. Recently, lower HbA1C values in the intensive arm were reported to be associated with reduced CV disease events [54] and difficulty in lowering HbA1C below 7% in the intensive group may actually have been the best predictor of CV events [54]. None of the trials in patients with longer duration of disease demonstrated a clear adverse (or beneficial) effect of any particular glucose-lowering medication, although the trials were not designed to answer this question. However, a reasonable interpretation is that there may be small to modest benefits on CV disease events with glucose lowering as long as such glucose lowering can be achieved without increased risk for hypoglycemia. In addition, the difficulty in understanding the results of the recent CV trials may be due, in part, to the declining incidence rate of CV events in patients with diabetes. The declining event rates observed in the most recent CV studies may suggest that improvements to CV risk may be due to other temporal phenomena. Resolving the problem of residual risk is increasingly difficult to prove experimentally because the low absolute risk of an event coupled with weight loss in clinical and observational studies of patients with type 2 diabetes [62,63]. Importantly, exenatide therapy does not increase the risk of severe hypoglycemia when used in the absence of agents commonly associated with hypoglycemia, primarily SFU and insulin. In several small studies, insulin-induced hypoglycemia was associated with lengthened QT intervals, which may increase the risk of arrhythmia [64,65]. It is possible that the low risk of hypoglycemia with GLP-1 receptor agonist treatments may contribute to a lower risk of CV events.

Exenatide therapy has been associated with reductions in multiple cardiovascular risk factors. A pooled analysis of randomized controlled clinical trials of exenatide BID demonstrated a significant reduction in systolic blood pressure with exenatide therapy compared with placebo or insulin therapy [66]; a similar result was obtained in a recent randomized controlled clinical trial [67]. In open-label extension studies lasting up to 3 years, exenatide BID treatment resulted in improved glycemic control accompanied by moderate weight loss and improvements in BP, cholesterol levels, inflammatory markers, and insulin resistance for some patients [15-17]. Improvements in postprandial lipidemia associated with exenatide treatment were identified in a 1-year open-label study that compared the efficacies of exenatide and insulin glargine [68,69]. Compared with insulin-treatment, significant decreases in post-prandial triglycerides, free fatty acids, HDL-C, VLDL-C, and Apo B48 were observed with exenatide treatment. No between-group differences were found in postprandial total cholesterol, LDL-C, ApoA1, ApoA2, Apo B100, or ApoC3, although exenatide reduced the post-prandial oxidative stress markers P-malondialdehyde and oxidized LDL [68]. Reductions in high-sensitivity C-reactive protein and increases in total adiponectin were also observed [69]. The latter changes were not statistically dependent on changes in fat mass or body weight [69].

Exenatide therapy has been associated with 1 or 2 bpm increases in heart rate in individual clinical trials [29,67]. In this analysis, mean heart rate changes were +0.5 bpm for exenatide and +0.1 bpm for the pooled comparators. The clinical importance of small increases in heart rate is unclear. The majority of epidemiological studies in patients with diabetes have studied the effects of 10 bpm increases in heart rate from any cause, which are associated with increased CV risk [70,71].

No effects of exenatide BID on cardiac repolarization (QT/QTc interval) have been observed in preclinical toxicology studies or during clinical studies of exenatide BID in patients with type 2 diabetes. A thorough QT study in healthy volunteers demonstrated that,
compared with placebo treatment, exenatide BID treatment was not associated with clinically significant QTc prolongation [72]. A thorough QT study for an investigational, long-acting formulation of exenatide is ongoing.

Extrapolating from pre-clinical studies, it is possible that GLP-1 receptors located in the heart and vasculature may play a protective role with respect to CV disease [73,74]. In animal models, GLP-1 reduced infarct size after coronary artery ischemia [75], improved left ventricular ejection volume in heart failure [73], improved glucose uptake in the myocardium [74], and induced nitric oxide-independent vasorelaxation in the endothelium [74]. Recent studies evaluated treatment with a GLP-1 receptor agonist in rodent models of severe hypertension or congestive heart failure. Dahl salt-sensitive rats fed a high-salt diet were found to have less hypertension, renal dysfunction, and mortality after 4 weeks of continuous therapy with a GLP-1 analog than were those without treatment [76]. Similarly, rats with heart failure after coronary artery ligation that were treated with a GLP-1 analog demonstrated improved cardiac function, cardiac dimension, exercise capacity, and survival compared to untreated rats [77]. In humans, there is limited evidence that GLP-1 improves left ventricular ejection fraction, reperfusion, or functional status in patients with heart failure or MI [78,79]. Further studies are warranted in patients with specific types of CV disease to understand if the latter effects and those on the endothelium are seen in humans. Studies of the CV effects of different formulations of GLP-1 receptor agonists, including an oral formulation, are also justified to determine whether similar results are seen in all members of this drug class [80].

Limitations
Major limitations of this analysis were the inclusion of studies of short duration, lack of complete data on CV history, and the lack of pre-specified and blinded adjudication of the CV events. In addition, the incidence of CV events was low in the individual trials. As with all meta-analyses, the current meta-analysis was retrospective in nature. Pooling the placebo group with a single active-comparator group is another potential limitation of this analysis. The rationale for pooling the placebo-and active-comparator patients was to provide greater power for this analysis by increasing the sample size and the number of observed outcomes.

Conclusions
On the basis of this integrated analysis of 3,945 participants representing 1,070 patient-years of exenatide exposure and 780 patient-years of PC exposure, no increase in CV risk was associated with use of exenatide BID (vs. PC) in patients with type 2 diabetes participating in clinical trials. This analysis, conducted in compliance with FDA guidance on the subject, provides insight into the CV safety of exenatide treatment.

Acknowledgements and Funding
We thank Carl Brown for earlier contributions to this analysis and Mary Beth DeYoung and Lee Meller for assistance with final revisions. We also thank Haiying Dong and the Amylin Statistical Programming Group for their contributions to this analysis. CB, MBDY, LM and HD are employees and stockholders of Amylin Pharmaceuticals, Inc. Dawn Nicewariner and Irina Yushmanova, MD were employees of Amylin Pharmaceuticals, Inc. at the time of this analysis.

This analysis was sponsored by Amylin Pharmaceuticals, Inc. and Eli Lilly and Co., which have a global agreement to collaborate on the development and commercialization of exenatide.

Authors’ contributions
RR reviewed and interpreted the data, requested additional analysis, and critically reviewed the manuscript; JH participated in the acquisition and statistical analysis of the data and reviewed the manuscript; DN participated in analysis and interpretation of the data and drafted the manuscript; IF participated in design of the analysis, data interpretation, and critical review of the manuscript; BJH participated in analysis and interpretation of the data, requested additional analysis, and drafted sections of the manuscript; LS contributed to the conception and design of the analysis, interpreted the analyzed data, and critically reviewed the manuscript. All authors read and approved the final manuscript.

Competing interests
RR has received research support during the last 5 years from Amylin Pharmaceuticals, Inc., Eli Lilly & Company, NovoNordisk, GSK, Roche, and Sanofi Aventis; he is not an employee or stock-holder of any of these companies and does not have any other financial competing interests. JH, DN, IF, and LS were employees and stockholders of Amylin Pharmaceuticals, Inc., at the time of this analysis. BJH is an employee and stockholder of Eli Lilly & Co. Amylin Pharmaceuticals, Inc. and holds the patent for synthetic Exendin-4 (exenatide).

Received: 5 November 2010 Accepted: 16 March 2011
Published: 16 March 2011

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